
Bayesian Model Averaging and some applications

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- Some recommendations and open questions

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- Some recommendations and open questions
- Key references

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- Model averaging: our inference is averaged over all the models in the model space considered, using weights that are either derived from Bayes’ theorem (BMA) or from sampling-theoretic optimality considerations (FMA). Here focus on BMA

BMA

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E.g. wish to predict the unobserved y_f on the basis of the observed y . Sampling model for y_f and y jointly is $p(y_f|y, \theta_j, M_j)p(y|\theta_j, M_j)$, where M_j is the model selected from K possible models, and $\theta_j \in \Theta_j$ are the parameters of M_j .

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Assign a (continuous) prior $p(\theta_j|M_j)$ for the parameters and a discrete prior $P(M_j)$ on the model space. Predictive distribution is

$$p(y_f|y) = \sum_{j=1}^K \left[\int_{\Theta_j} p(y_f|y, \theta_j, M_j)p(\theta_j|y, M_j)d\theta_j \right] P(M_j|y) \quad (1)$$

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Averaging at two levels: over parameter values, given each possible model, and discrete averaging over all possible models

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Square brackets in (1): predictive given M_j obtained using the posterior of θ_j given M_j , which is

$$p(\theta_j|y, M_j) = \frac{p(y|\theta_j, M_j)p(\theta_j|M_j)}{\int_{\Theta_j} p(y|\theta_j, M_j)p(\theta_j|M_j)d\theta_j} \equiv \frac{p(y|\theta_j, M_j)p(\theta_j|M_j)}{p(y|M_j)}, \quad (2)$$

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Denominators of both averaging operations are made explicit in (2) and (3). $p(y|M_j)$ in (2) is the marginal likelihood of M_j and is a key quantity: Bayes factor is the ratio of marginal likelihoods (posterior odds = Bayes factor * prior odds). $p(y)$ in (3) is a sum (challenge often lies in the number of models K).

More generally, the posterior distribution of any quantity of interest, say Δ , which has a common interpretation across models is a mixture of the model-specific posteriors with the posterior model probabilities as weights, *i.e.*

$$P_{\Delta|y} = \sum_{j=1}^K P_{\Delta|y, M_j} P(M_j | y). \quad (4)$$

Construction of the Model Space

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Three key sources of uncertainty:

- Theory: which variables are important drivers? Often, theories regarding variable inclusion do not contradict each other (“open-endedness” of the theory)
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Usually theoretical results are derived under the assumption that \mathcal{M} contains the “true” data-generating model (“ \mathcal{M} -closed”), but most important results like model selection consistency extend to “ \mathcal{M} -open” settings in an intuitive manner.

Covariate uncertainty in normal linear regression

Most common setting: model uncertainty about which covariates to include, *i.e.* under model j the n obs. in y are generated from

$$y|\theta_j, M_j \sim N(\alpha\iota + Z_j\beta_j, \sigma^2). \quad (5)$$

Here ι is a vector of ones, Z_j groups k_j of the possible k regressors and $\beta_j \in \Re^{k_j}$ are the regression coefficients. All models contain an intercept $\alpha \in \Re$ and a scale $\sigma > 0$ with a common interpretation.

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Economics: k up to 100 (growth), so $K > 10^{30}$ and need efficient computational tools. Genetics (usually $n \ll k$): k could be up to 100,000, leading to $K > 10^{30,000}$!

BMA: Prior structures

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Prior on model parameters:

$$p(\alpha, \beta_j, \sigma \mid M_j) \propto \sigma^{-1} f_N^{kj}(\beta_j \mid 0, \sigma^2 g(Z_j' Z_j)^{-1}), \quad (6)$$

which is a “*g*-prior” and leads to closed form for integral in (1) and the marginal likelihood (likelihood integrated out with the prior).

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Only requires choice of two scalars g and w : hyperpriors are recommended (adaptive and more robust)

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- BMA predicts at least as well as any single model (assuming data is generated by (1)) and there is ample empirical evidence for clear superiority (probabilistic forecasts)
- BMA leads to point estimates that minimize MSE and BMA estimation intervals are calibrated in the sense that the average coverage probability of a BMA interval with posterior probability α is at least equal to α (assuming data is generated by (1)).

Numerical methods for large model spaces

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So no need to avoid using large \mathcal{M} !

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Posterior odds between models, given g and w :

$$\frac{P(M_i | y, w, g)}{P(M_j | y, w, g)} = \left(\frac{w}{1-w} \right)^{k_i - k_j} (1+g)^{\frac{k_j - k_i}{2}} \left(\frac{1 + g(1 - R_i^2)}{1 + g(1 - R_j^2)} \right)^{-\frac{n-1}{2}} \quad (8)$$

The three factors correspond to a model size penalty induced by the prior on the model space, a model size penalty resulting from the marginal likelihood and a lack-of-fit penalty from the marginal likelihood.

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Hyperpriors on g and w can have a large effect on the induced penalties for model complexity but not on the impact of the relative fit of the models

Complexity penalties

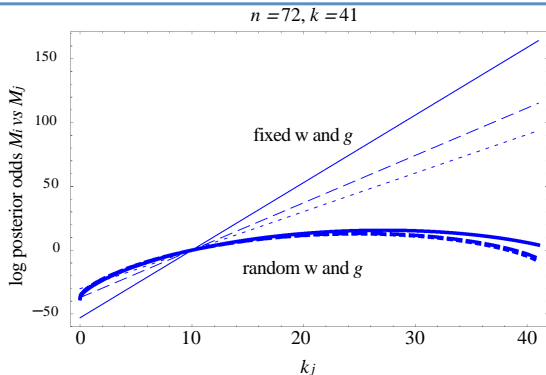


Figure: Posterior odds as a function of k_j when $k_i = 10$ with equal fit, using prior mean model size $m = 7$ (solid), $m = k/2$ (dashed), and $m = 2k/3$ (dotted). Bold lines correspond to random w and g

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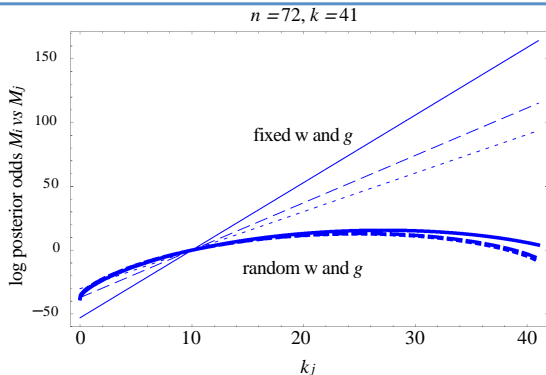


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Hyperpriors more robust to m , less extreme and penalize models of size around $k/2$ (multiplicity penalty)

Approximations and hybrids

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Hybrids of frequentist and Bayesian methods were used e.g. to deal with endogenous regressors: BIC approximations to posterior model probabilities for averaging over classical two-stage least squares (2SLS) estimates.

Other sampling models

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- Generalized linear models (GLMs), for example logistic, probit or ordered response models
- Generalized additive models (nonlinear effects)
- Models for outliers and non-normal models (e.g. Student- t)
- Dynamic models, e.g. AR(F)IMA and DLMs
- Models with endogenous regressors (IV models)
- Models for longitudinal data with individual effects
- Models for spatial data (Spatial AR models)
- Duration models

Endogeneity

Endogeneity occurs if one or more of the covariates is correlated with the error term in the equation corresponding to (5). In particular, consider the following extension of the model in (5):

$$y = \alpha\iota + x\gamma + Z_j\beta_j + \varepsilon \quad (9)$$

$$x = W\delta + \nu, \quad (10)$$

where x is an endogenous regressor and W is a set of instruments, independent of ε . The error terms are iid:

$$(\varepsilon_i, \nu_i)' \sim N(0, \Sigma), \quad (11)$$

with $\Sigma = (\sigma_{ij})$ a 2×2 covariance matrix. Whenever $\sigma_{12} \neq 0$ this introduces a bias in the OLS estimator of γ and a standard classical approach is the use of 2SLS estimators instead.

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Lenkoski et al. (2014) also account for model uncertainty in the selection of instruments. They propose a two-step procedure that first averages across the first-stage models (for the endogenous variables) and then, given the fitted endogenous regressors from the first stage, it again takes averages in the second stage. Both steps use BIC weights (approximations).

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Karl and Lenkoski (2012) propose IVBMA, which uses conditional Bayes factors to account for model uncertainty within a Gibbs algorithm. Their algorithm hinges on certain restrictions (e.g. joint Normality and conditionally conjugate priors), but it is exact, efficient and is implemented in an R-package.

Application in Economics

Lenkoski et al. (2014). *Econometric Reviews*

Three main theories about what drives economic growth:
geography (natural and human resources), international trade
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In previous literature some (influential) studies found evidence that property rights are a strong driver for growth, but without considering many alternative models. Similarly, others concluded that trade variables were key drivers (without controlling for the effect of institutions). Rodrik et al. (2004) (RST) provide a “horse race” among alternative theories that propose candidate instruments and regressors, but don’t use BMA (they just compare a limited set of models) and conclude only institutions matter

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Consider e.g. Rule of Law and Integration (Openness):

Models	Rule of Law			Integration		
	PIP	mean	sd	PIP	mean	sd
RST core	1.00	1.28	0.18	0.20	0.11	0.26
limited \mathcal{M}	1.00	0.95	0.13	0.07	0.07	0.14
full \mathcal{M}	0.96	0.80	0.32	0.85	0.93	0.38

Table: Some BMA (2nd stage) results with different sets of possible covariates (PIP is posterior inclusion probability)

Divergence of results (between 2SLS and BMA) grows as we allow for more uncertainty (bigger model spaces). Integration becomes important driver and all three theories are supported in the BMA results using all available variables.

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Hypotheses: larger district magnitude (more proportionality) could make potential voters more likely to

- feel represented
- feel close to a political party
- be contacted by a political party
- turn out to vote

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In contrast with some of the literature, but this study takes into account model uncertainty in a reasonably large model space (with other potential explanatory variables) rather than focus on a very limited set of models

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Rainey, Electoral Studies (2016)

BMA explaining survey outcomes addressing the hypotheses through 16 potential regressors (logistic regression with BIC approximation to compute posterior model probabilities)

Posterior probability that district magnitude has no effect on each of the four possible outcomes is at least 0.97

In contrast with some of the literature, but this study takes into account model uncertainty in a reasonably large model space (with other potential explanatory variables) rather than focus on a very limited set of models

Important predictors for voting turnout are age and marital status

Software and resources

A number of free R-packages: BMS, BAS, BayesVarSel, ivbma
(endogenous regressors)

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Some useful resources available online, e.g.

<http://bms.zeugner.eu/resources/> (also introductory material)

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- Lots of unexplored areas especially outside of the normal linear model (prior structures, elicitation and effects; properties; computation)
- I hope that BMA can become a key methodology in many areas of application (it already is in macroeconomics), and can contribute to constructive communication by better understanding the reasons for differences in empirical findings

Some useful references

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